GEOMETRIC STRUCTURES ON A STATISTICAL MANIFOLD AND GEOMETRY OF ESTIMATION

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by

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ABSTRACT

The main objective of this thesis is to study the various geometric structures on a statistical manifold and the geometry of parameter estimation. This study comes under the area of Information Geometry which is the geometric study of a statistical model of probability distributions. A statistical model equipped with a Riemannian metric and a pair of dual affine connections is called a statistical manifold. Amari's α -geometry is an important geometric structure on a statistical manifold which plays a major role in the asymptotic theory of estimation.

In Chapter 2 we introduce a generalized class of geometric structures on a statistical manifold called the (F, G)-geometry using a general embedding function F and a positive smooth function G. In Section 2.2 the Fisher information metric and the α -connections are computed for a statistical manifold defined on finite sets. In Theorem 2.3.5 we prove a necessary and sufficient condition for two (F, G)-connections to be dual with respect to the G-metric. In Theorem 2.3.6 we show that the α -geometry is a special case of the (F, G)-geometry. Thus we obtain a generalized dualistic structure on a statistical manifold which includes the α -geometry as a special case. Further the G-metric and the (F, G)-connections are computed for statistical manifold defined on finite sets in Section 2.3.

In Chapter 3 we study the invariance properties of various geometric structures on a statistical manifold and classify them into invariant and non-invariant classes. The covariance under reparametrization of the (F, G)-geometric structures are shown in Theorems 3.2.3 and 3.2.4. Then in Theorem 3.2.5 we prove that the (F, G)-geometry is not invariant under smooth one to one transformations of the random variable in general. In Corollary 3.2.6 we prove that the α -geometry is the only (F, G)-geometry which is invariant under smooth one to one transformations of the random variable. In Theorems 3.2.7 and 3.2.8 we show that the (α, ρ, τ) -geometry is covariant under reparametrization and is not invariant under smooth one to one transformations of the random variable. In South and the random variable in general. Also the α -geometry is the only (α, ρ, τ) -geometry which is invariant under smooth one to one transformations of the random variable in general. the (F,G) and (α, ρ, τ) -geometries are given in Theorem 3.2.11.

In Chapter 4 first we give the (± 1) -conformal equivalence of the α -geometry and the geometry induced from the conformal transformation of the α -divergence in Propositions 4.2.3 and 4.2.4. In Corollary 4.2.6 we prove that the q-structure is the conformal flattening of the α -geometry. Then we discuss the importance of non-invariant (F, G)-geometry in the study of the dually flat geometries of the deformed exponential family. There are two dually flat geometries on a deformed exponential family, the Ugeometry and the χ -geometry. In Theorem 4.3.4 we show that the U-geometry is the (F, G)-geometry for suitable choices of F and G. Further we prove that the χ -geometry is the conformal flattening of the (F, G)-geometry for suitable choices of F and G in Theorems 4.3.16, 4.3.17 and 4.3.18.

In Chapter 5 we consider the parameter estimation problem based on a mismatched model. In Theorems 5.3.1 and 5.3.2 we prove a necessary and sufficient condition for the estimator based on a mismatched model to be consistent and first order efficient. Further a theoretical formulation of the maximum likelihood estimation problem based on a mismatched model in an exponential family is given. We prove a necessary and sufficient condition for an MLE based on a mismatched model to be consistent and efficient in Theorems 5.3.8 and 5.3.9.

In Chapter 6 we define certain generalized notions like *F*-product, *F*- independence of random variables and maximum *F*-likelihood estimator (*F*-MLE) in Section 6.1. In Theorem 6.1.6 we show that the *F*-MLE is a MAP estimator with a prior. Then using the *F*-escort probability distribution we define two generalized notions of MLE, the x_N -based *F*-escort MLE and the *F*-escort MLE based on the product of *F*-escort distribution of the marginal probability density of single observations in Section 6.2. In Theorem 6.2.3 we give a characterization of the *q*-escort MLE among the x_N based *F*-escort MLE as a Bayesian MAP estimator with a prior. Further an analytic proof of the *F*-version of the maximum entropy theorem is given in Theorem 6.2.5. In Theorem 6.3.2 a proof of the generalized Cramer-Rao bound defined by Naudts is given. Further we show that the *U*-estimator for the dual coordinate in the *U*-geometry of the deformed exponential family is optimal with respect to this bound in Theorem 6.3.3. This chapter ends with an open problem regarding the properties of the *F*-MLE in a deformed exponential family.